

# Modeling multivariate data: from precipitation to finance

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## Abstract:

This talk will outline my ongoing and future planned projects, discussing possibilities for students' research. Two examples: spatial random field modeling (based on the MS work with Anandkumar Shetiya), and a future project with stochastic volatility modeling.

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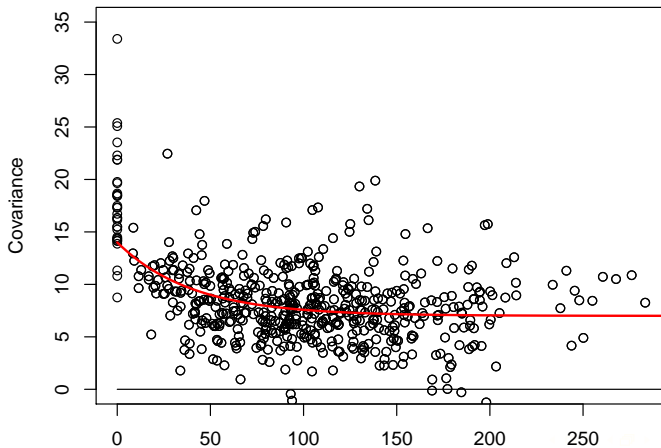
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- ▶ Examples: precipitation measurements at gauge  $j$ , asset prices (asset =  $j$ , time =  $t$  discretized)

Consider purely spatial dependency in precipitation.

Typically, if we plot covariance of precipitation measurements versus distance we observe something like this



# Modeling spatial dependency

Need to model the dependency of  $cov(Y^{(i)}, Y^{(j)})$  on the distance.

- ▶ • Traditional geostatistical methods based on variogram modeling: basically, model some function  $\phi$ ,

$$cov(Y^{(i)}, Y^{(j)}) = \phi(\mathbf{r}_i - \mathbf{r}_j)$$

$\mathbf{r}_j$  is the location of gauge  $j$

(Math 586, Spring '09)



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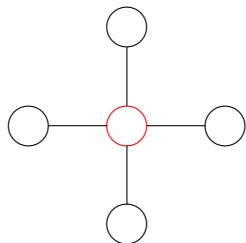
- ▶ • Convolution of GMRF's (Gaussian Markov random fields), which are easier to deal with computationally.

## GMRF: Some basics

Markov field: (extension from 1-dim concept of Markov chain)  
possesses *Markov property*

$$P(X(\mathbf{s}_i) | \text{rest of } X) = P(X(\mathbf{s}_i) | \{X(\mathbf{s}_j), j \sim i\})$$

i.e. the value at the location  $i$  depends only on the values at *neighboring* locations  $j \sim i$ . E.g. rectangular grid neighbors



## GMRF: Some basics

*Gaussian* vector  $\mathbf{x} = \{X(s_j), j = 1, \dots, m\}$  is determined by its prior distribution

$$p(\mathbf{x}) \propto \exp(-\lambda_x \mathbf{x}^T V \mathbf{x}) \quad (1)$$

with *precision matrix*  $V$  defined by

$$V_{ij} = \begin{cases} -1, & i \sim j \\ \text{number of neighbors of } i, & i = j \\ 0, & \text{otherwise} \end{cases}$$

and  $\lambda_x$  is the smoothness parameter.

Advantage: matrix  $V$  is sparse.

# GMRF: convolution

However, the gauges are not on a regular grid!

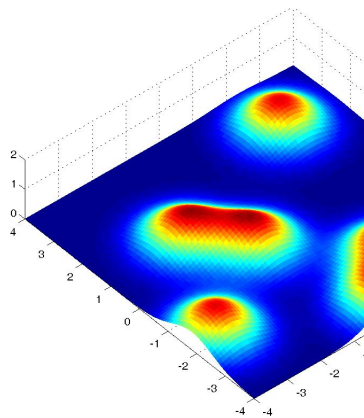
To fill in the values, use **convolutions**,  
at location  $\mathbf{s}$

$$\tilde{X}(\mathbf{s}) = \sum_j X(\mathbf{s}_j) \psi(\mathbf{s} - \mathbf{s}_j)$$

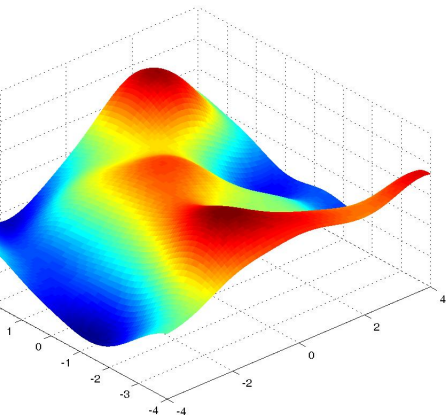
where  $\psi$  is some kernel function, say  
Gaussian

$$\psi(\mathbf{s}) = \exp(-\mathbf{s}^T \mathbf{s} / 2\sigma_\psi^2),$$

and  $\sigma_\psi$  can be taken equal to the grid  
spacing.



# GMRF: convolution



Finally, our observations

$$Y(\mathbf{s}) = \sum_{j \in \text{grid}} X(\mathbf{s}_j) \psi(\mathbf{s} - \mathbf{s}_j) + e(\mathbf{s}) \quad (2)$$

where  $e$  are white-noise (i.i.d.) errors

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- ▶ Handy tool for inverting: Bayesian approach.  
Given  $p(\mathbf{x}, \theta) =$  prior, and  $p(\mathbf{y}|\mathbf{x}, \theta) =$  likelihood, our goal is to compute

$$\text{Posterior: } p(\mathbf{x}, \theta|\mathbf{y}) \propto \text{prior} \cdot \text{likelihood}$$

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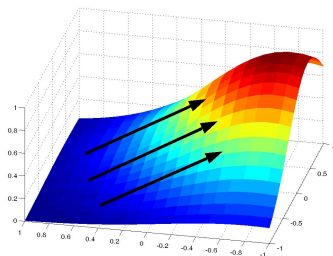
- ▶ Computational tools: Gibbs sampler.  
End up with a Markov Chain Monte Carlo (MCMC) sample.



# MFD random field

Effects of terrain: **MFD (moisture flux direction)**

Increased precipitation when moisture is moving upslope.



Compute another random field  $W$  that helps in improving the prediction.

$W$  is directional, i.e. its values are between  $0$  and  $2\pi$

## MFD random field: prior

Try to follow the convolution approach.

Let vector  $W$  = the values of MFD at the grid points.  
We set a prior on  $W$  that encourages it to be smooth:

$$p(W) \propto \exp \left[ \gamma \sum_{i \sim j} \cos(W_i - W_j) \right]$$

The values of  $W$  that are close (in circular topology) will receive a high prior weight, and diametrically opposite values of  $W$  receive a low prior weight.

$\gamma \geq 0$  is a smoothness parameter.

# MFD random field: convolution

To find  $W$  co-located with the data  $Y$ , follow the convolution

$$\tilde{W}(\mathbf{s}) = \circ \sum_{j \in \text{grid}} \psi(\mathbf{s} - \mathbf{s}_j) W_j$$

however the sum taken above is understood in the circular topology, that is, through weighted averages

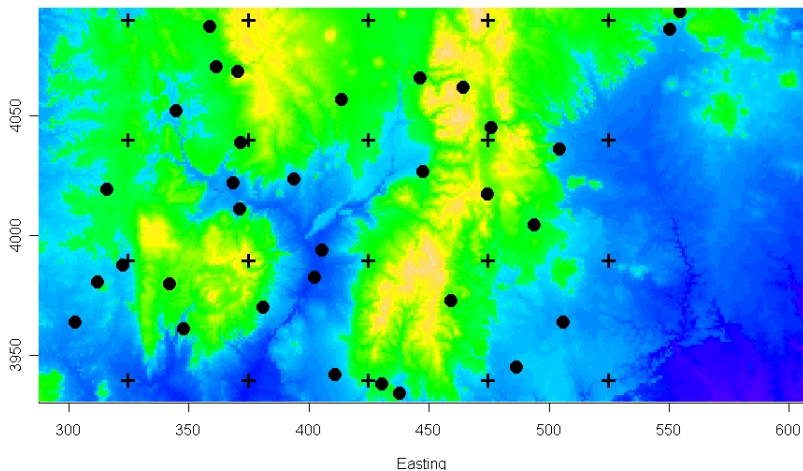
$$\sin(\tilde{W}(\mathbf{s})) = \frac{\sum_j \psi(\mathbf{s} - \mathbf{s}_j) \sin(W_j)}{\sum_j \psi(\mathbf{s} - \mathbf{s}_j)}, \quad \cos(\tilde{W}(\mathbf{s})) = \dots$$

The overall model becomes

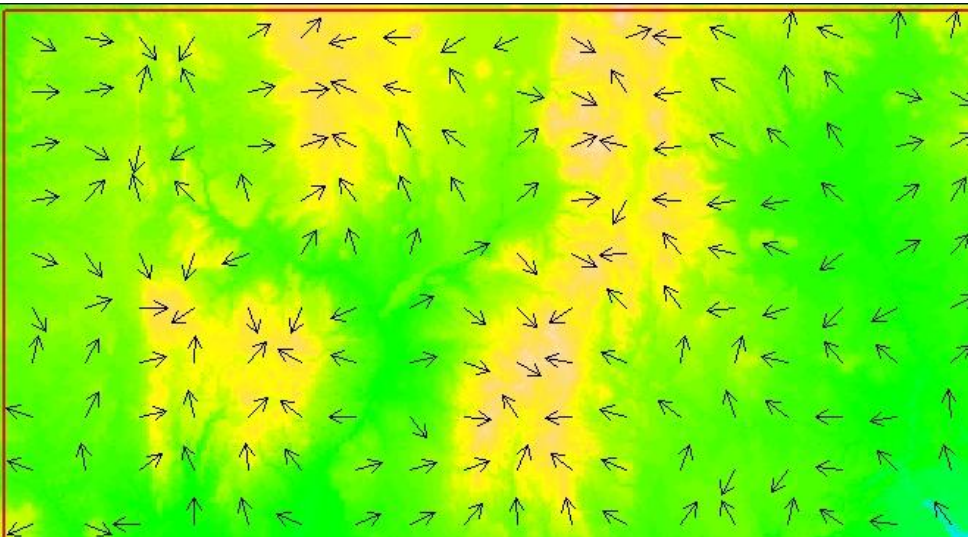
$$Y(\mathbf{s}) = \sum_j X(\mathbf{s}_j) \psi(\mathbf{s} - \mathbf{s}_j) + \beta \cos(\tilde{W}(\mathbf{s}) - A(\mathbf{s})) + e(\mathbf{s}) \quad (3)$$

$A(\mathbf{s}) =$  terrain aspect at location  $\mathbf{s}$

# Study area

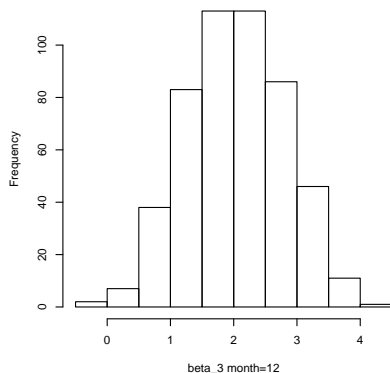
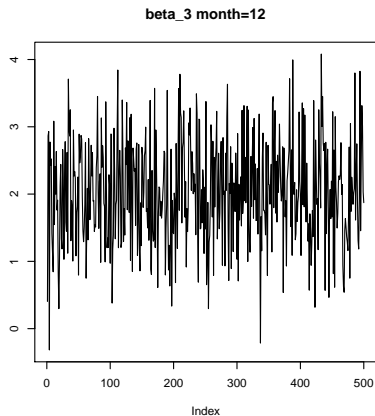


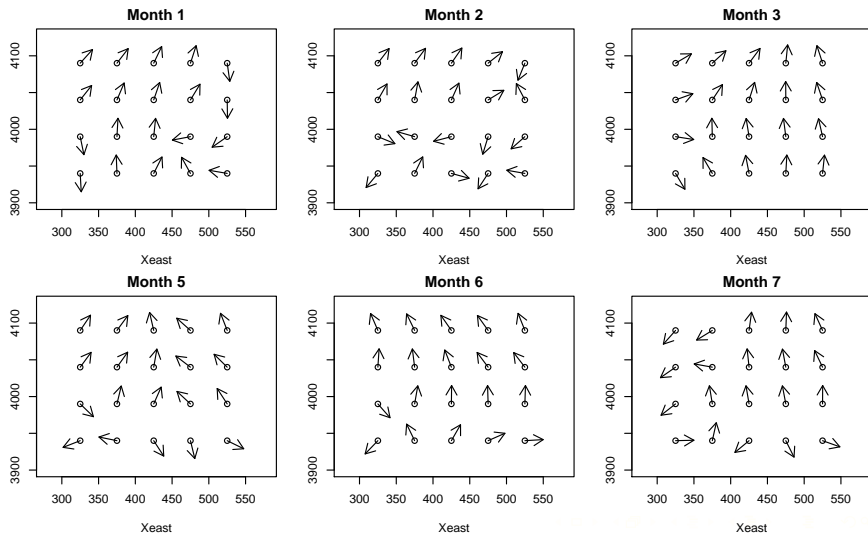
# Study area: aspects

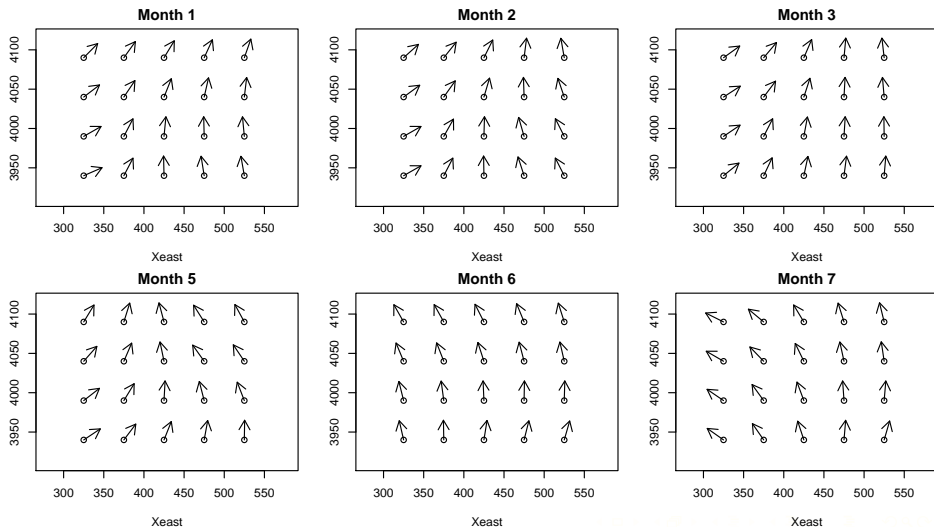


# Output: $\beta$

This shows a typical MCMC output for the posterior distribution of  $\beta$ .



Fitted MFD,  $\gamma = 2$ 

Fitted MFD,  $\gamma = 5$ 



# Modeling financial data

# Asset price model

Model the asset price  $X_t$  as a random walk

$$X_t = X_{t-1} + e_t$$

$\{e_t\}$  are independent

However, they are not identically distributed:

$$e_t \sim \text{Normal}(0, \sigma_t)$$

where  $\sigma_t =$  “volatility” is itself time-varying; and it is predictable to some extent.

Dow



# Stochastic volatility

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This is an autoregressive model

$$\log(\sigma_t) = \mu + r(\log(\sigma_{t-1}) - \mu) + \nu_t$$

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**Research question:** for different assets  $j$ , does the clustering of  $\sigma_t^{(j)}$  take place?

That is, if we look at multiple assets, do they form the groups within which  $\sigma_t^{(j)}$  behave alike?

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- ▶ Multivariate Normal distribution
- ▶ Time-series approaches, autocorrelation, autoregression etc.
- ▶ Bayesian inference, computational methods (Markov Chain Monte Carlo)



QUESTIONS?

THANK YOU!

`www.nmt.edu/~olegm/talks/Multivar/`  
for pdf file