

# Lecture 2: Probability: many variables

Math 586

**Recap:** (Lecture 1)

- “A *random variable* is a variable whose values are randomly generated according to some probabilistic mechanism” - Isaaks & Srivastava
- $X$  = random variable,  $x$  = number.
- $F(x) = P(X \leq x) = \int_{-\infty}^x f(t) dt$  (cumulative) distribution function.
- $f(x)$  density (continuous) or  $P(X = x_i)$  (discrete)

## Joint distribution

- If  $X_1, \dots, X_n$  are r.v. then

$$F(x_1, \dots, x_n) = P(X_1 \leq x_1, X_2 \leq x_2, \dots, X_n \leq x_n)$$

is their **joint distribution function**.

- If  $F(x_1, \dots, x_n)$  is differentiable in each  $x_i$  then

$$f(x_1, x_2, \dots, x_n) = \frac{\partial^n F(x_1, x_2, \dots, x_n)}{\partial x_1 \partial x_2 \dots \partial x_n}$$

is their **joint density function**.

- If  $\mathbf{X} = (X_1, X_2, \dots, X_n)'$  is a **random vector** (column vector,  $n \times 1$ ), and subset  $A \subset \mathbb{R}^n$  then

$$P(\mathbf{X} \in A) = \int \int_A \dots \int f(x_1, x_2, \dots, x_n) dx_1 dx_2 \dots dx_n$$

- Expectation of a function

$$\mathbb{E}[g(X_1, \dots, X_n)] = \int \int \dots \int g(x_1, x_2, \dots, x_n) \cdot f(x_1, x_2, \dots, x_n) dx_1 dx_2 \dots dx_n$$

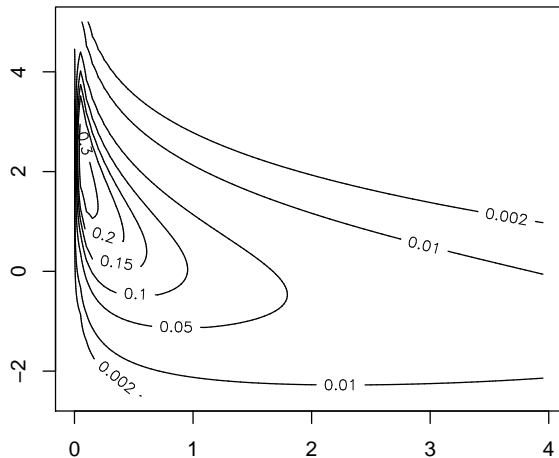
- Statistical independence

$X_1, \dots, X_n$  are **statistically independent** iff

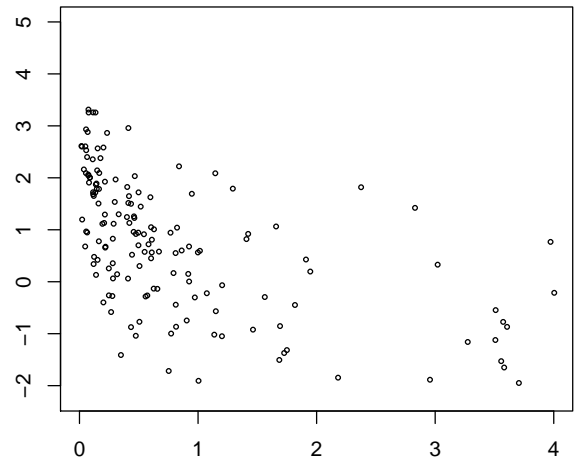
$$f(x_1, x_2, \dots, x_n) = f_1(x_1) \cdot f_2(x_2) \cdot \dots \cdot f_n(x_n)$$

One or more r.v.'s can be functionally dependent even though they are statistically independent.

**Contour plot of some bivariate distribution**



**Scatter plot of 1000 samples**



## Estimates of Distributions

Conceptual model: Population (of all feasible observations) from which we draw samples to estimate distribution and its properties, such as expected value.

**Example:** consider a manufactured item with design engineering strength, but actual strength varies in production. The “model strength” is a r.v.  $X$  with distribution  $F(x)$ . We don't know  $F$  *a priori* but must estimate it from the data.

- Estimate  $F(x_0)$  based on  $n$  samples  $x_1, x_2, \dots, x_n$ , e.g. using empirical CDF

$$\hat{F}(x_0) = \frac{\#\{x_i \leq x_0\}}{n} \quad \text{a.k.a. } \textit{ogive}$$

- Estimate the mean strength by using sample mean

$$\bar{x} = \frac{1}{n} \sum_{i=1}^n x_i$$

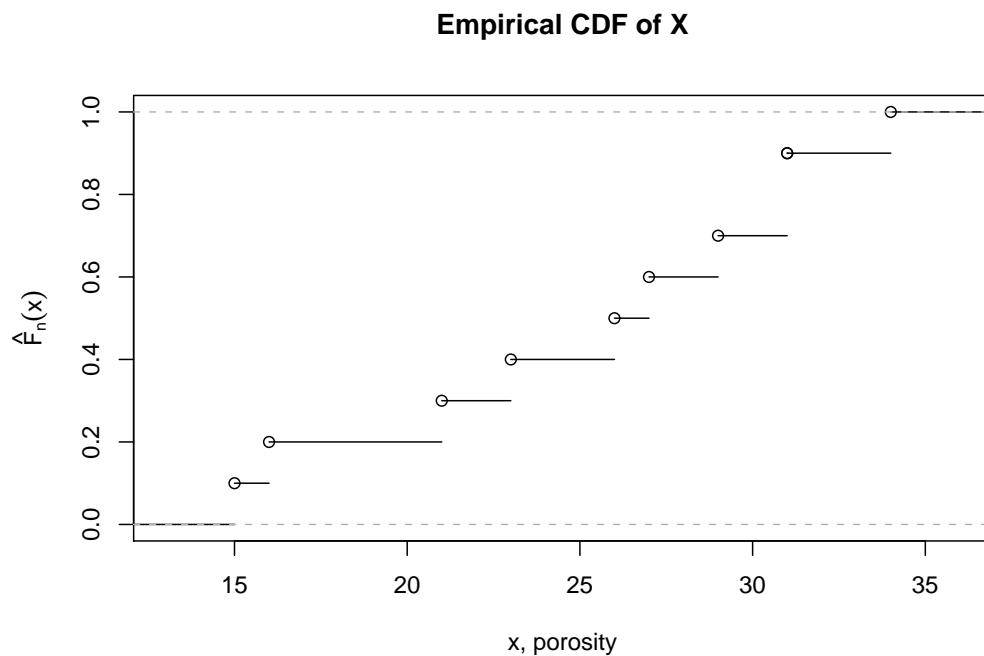
This is an estimator for  $\mathbb{E}[X]$ .

- Note: there is an important difference between the estimate ( $\hat{F}$  or  $\bar{x}$ ) and the true value.
- Algorithm to calculate  $\hat{F}$ : sort  $x_i$ 's from smallest to largest, get  $x_1^*, x_2^*, \dots, x_n^*$  then

$$\hat{F}(x) = \frac{j}{n}, \quad x_j^* \leq x \leq x_{j+1}^*$$

example: take  $n = 10$  samples of porosity (in %):

34, 27, 15, 23, 21, 31, 26, 29, 16, 31    reorder:  $\Rightarrow$  15, 16, 21, 23, 26, 27, 29, 31, 31, 34



Also, calculate sample mean  $\bar{x}$ .

## Variance and Covariance

- **Variance**

- Let  $\mathbb{E}[X] = \mu$ .

$$\text{Var}(X) = \sigma^2 = \mathbb{E}[(X - \mu)^2] \quad \text{2nd central moment}$$

- $\sigma$  is the Standard Deviation

- Also  $\text{Var}(X) = \mathbb{E}[X^2 - 2\mu X + \mu^2] = \mathbb{E}[X^2] - 2\mu\mathbb{E}[X] + \mu^2 \Rightarrow$

$$\text{Var}(X) = \mathbb{E}[X^2] - \mu^2 \quad \text{“Computational formula for variance”}$$

- **Covariance**

- Given r.v.'s  $X_1$  and  $X_2$  with means  $\mu_1, \mu_2$ ,

$$Cov(X_1, X_2) = \mathbb{E}[(X_1 - \mu_1)(X_2 - \mu_2)] = \mathbb{E}[X_1 X_2] - \mu_1 \mu_2$$

Note: 
$$Cov(X, Y) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} (x - \mu_x)(y - \mu_y) f(x, y) dx dy$$

Replace integral by summation for discrete case.

- If  $X_1$  and  $X_2$  are statistically independent then

$$Cov(X_1, X_2) = 0$$

- **Correlation coefficient** between  $X_1$  and  $X_2$  with st.dev.  $\sigma_1, \sigma_2$

$$\rho = \frac{Cov(X_1, X_2)}{\sigma_1 \sigma_2}$$

- **Variance of the sum:**

$$\begin{aligned} Var(a_1 X_1 + a_2 X_2) &= \mathbb{E}[(a_1 X_1 - a_1 \mu_1 + a_2 X_2 - a_2 \mu_2)^2] = \\ &= a_1^2 \mathbb{E}[(X_1 - \mu_1)^2] + a_2^2 \mathbb{E}[(X_2 - \mu_2)^2] + 2a_1 a_2 \mathbb{E}[(X_1 - \mu_1)(X_2 - \mu_2)] = \\ &= a_1^2 Var(X_1) + a_2^2 Var(X_2) + 2a_1 a_2 Cov(X_1, X_2) \end{aligned}$$

Thus,

$$Var(a_1 X_1 + a_2 X_2) = a_1^2 Var(X_1) + a_2^2 Var(X_2) + 2a_1 a_2 Cov(X_1, X_2)$$

- If  $X_1, X_2$  are independent, then  $Cov = 0$  and

$$Var(a_1 X_1 + a_2 X_2) = a_1^2 Var(X_1) + a_2^2 Var(X_2)$$

- Consider  $n$  r.v.'s,  $X_1, \dots, X_n$  then

$$Var\left(\sum_{i=1}^n a_i X_i\right) = \sum_{i=1}^n a_i^2 Var(X_i) + \sum_i \sum_{j \neq i} a_i a_j Cov(X_i, X_j)$$

*Note:* kriging algorithms are based on minimizing variance of linear combinations of r.v.'s. This expression for variance is very important. The covariance on RHS will carry information about spatial continuity.

- If  $X_1, X_2, \dots, X_n$  are independent, then  $Var\left(\sum_{i=1}^n a_i X_i\right) = \sum_{i=1}^n a_i^2 Var(X_i)$